

Statistical Estimation of the Parameters of a PDE

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- Nomenclature for image recovery
- Statistical model for inverse problems
- Traditional approaches – deconvolution example
- Recovering electrical conductivity via inference

Image Recovery nomenclature for Inverse Problems

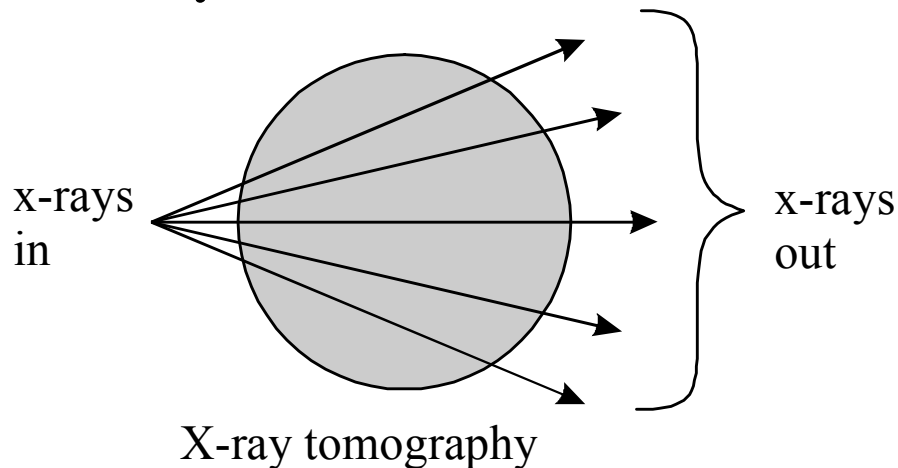


Image: spatially varying quantity of interest

optical reflectance of a scene

optical or radio brightness of sky

sound speed in tissue / ocean / earth

electrical conductivity of tissue / mud

Recovery: estimate image from indirect data

Forward Problem

image \longrightarrow data

physical model (PDE)

direct computation

well posed

unique

Inverse Problem

data \longrightarrow image

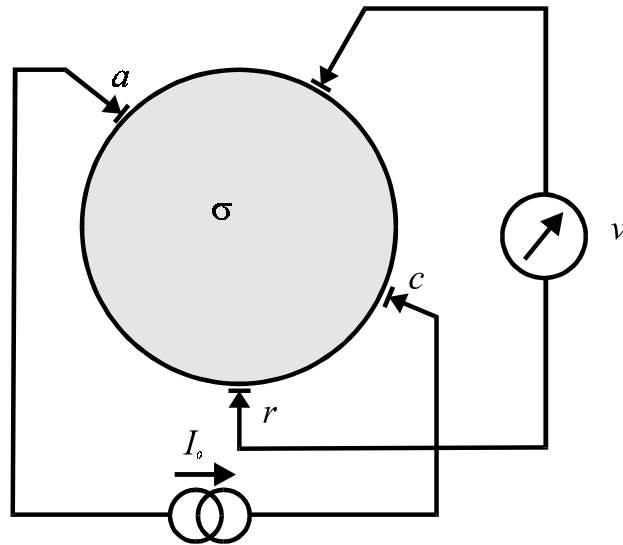
implicit

indirect

ill posed

never unique

Conductivity Imaging Measurement Set



- Electrodes at x_1, x_2, \dots, x_E
- Assert currents at electrodes $j = (j(x_1), j(x_2), \dots, j(x_E))^T$
- Measure voltages $v = (\phi(x_1), \phi(x_2), \dots, \phi(x_E))^T$.

Unknown $\sigma(x)$ related to measurements via Neumann BVP

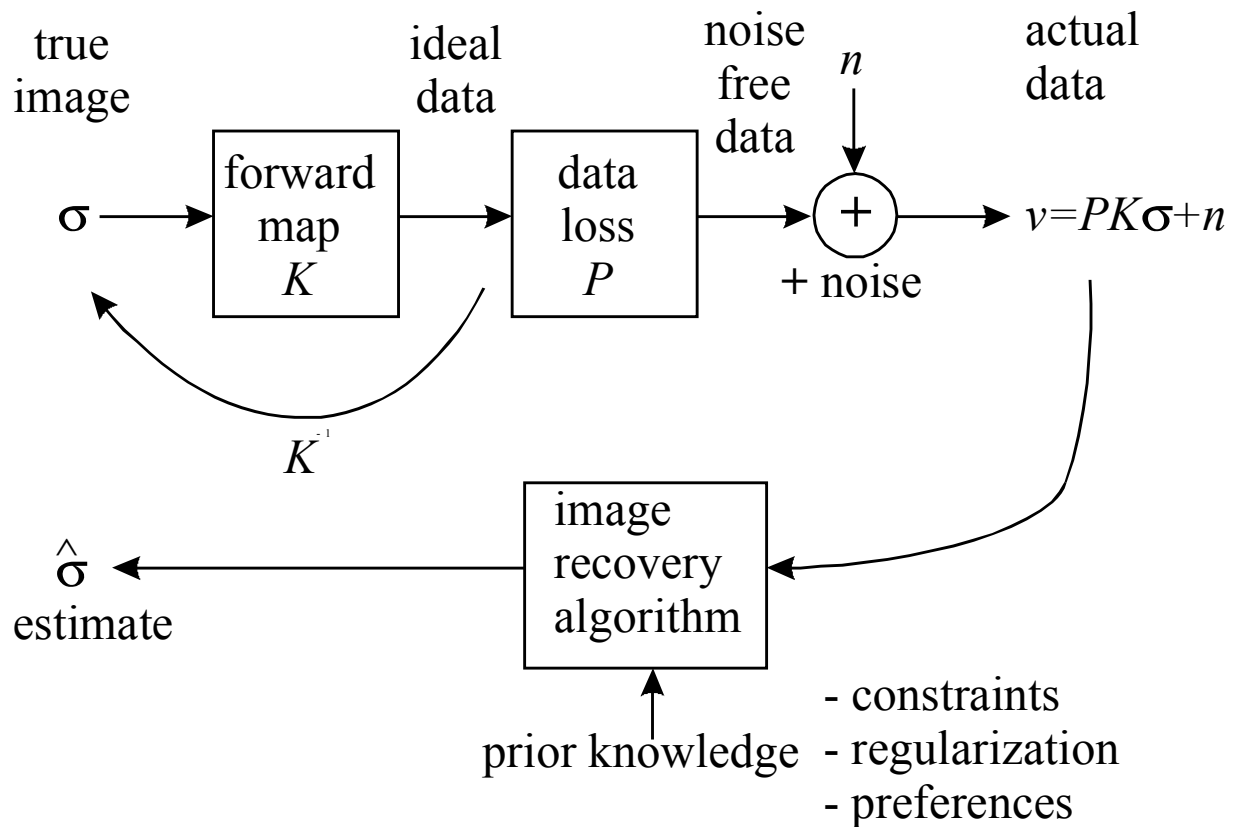
$$\begin{aligned} \nabla \cdot \sigma(x) \nabla \phi(x) &= 0 & x \in \Omega \\ \sigma(x) \frac{\partial \phi(x)}{\partial n(x)} &= j(x) & x \in \partial\Omega \end{aligned}$$

- Set of measurements is current-voltage pairs

$$\{j^n, v^n\}_{n=1}^N$$

Inverse problem is to find σ from these measurements
(non linear)

Statistical model of Inverse Problem



If $n \sim N(0, s^2)$, $v \sim N(PK\sigma, s^2)$

Given measurements v , the likelihood for σ is

$$L_v(\sigma) \equiv \Pr(v|\sigma) \propto \exp(-|v - \phi(\sigma)|^2 / 2s^2)$$

Posterior distribution for σ conditional on v

$$\Pr(\sigma|v) = \frac{\Pr(v|\sigma) \Pr(\sigma)}{\Pr(v)} \quad (\text{Bayes rule})$$

$\Pr(\sigma)$ is the prior distribution

Solutions = Summary Statistics

All information contained in posterior distribution $\Pr(\sigma|v)$

Traditional Solutions - modes

$$\hat{\sigma}_{\text{MLE}} = \arg \max L_v(\sigma) \equiv \arg \max \Pr(v|\sigma)$$

$$\hat{\sigma}_{\text{MAP}} = \arg \max \Pr(\sigma|v) \equiv \arg \max \Pr(v|\sigma) \Pr(\sigma)$$

e.g. simple Gaussian prior: $\Pr(\sigma) \propto \exp\left(-|\sigma|^2/2\lambda^2\right)$

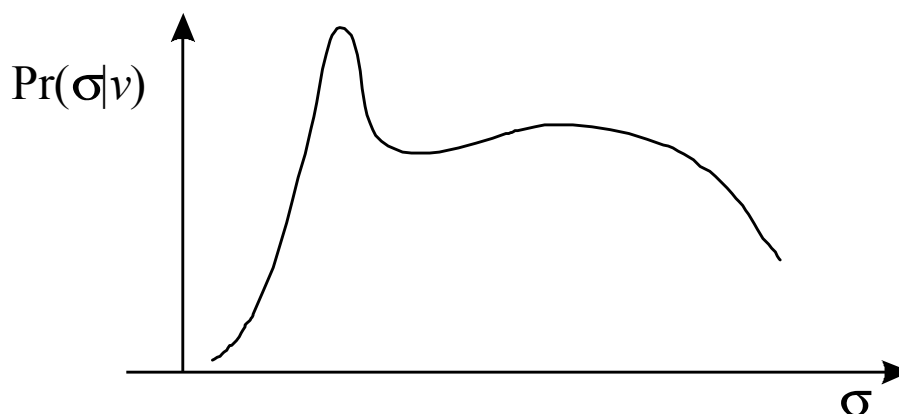
$$\hat{\sigma}_{\text{MAP}} = \arg \min |v - \phi(\sigma)|^2 + \alpha |\sigma|^2 \quad \alpha = s^2/\lambda^2$$

- Tikhonov regularization, Kalman filtering, Backus-Gilbert
- $\alpha \rightarrow 0$ Moore-Penrose inverse, $\alpha = 0$ least-squares

Inferential Solutions

“Answers” are expectations over the posterior

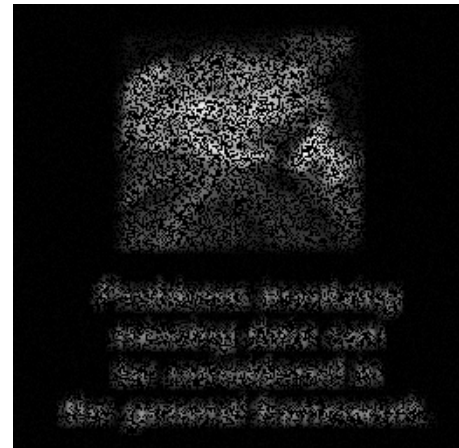
$$\mathbb{E}[f(\sigma)] = \int \Pr(\sigma|v) f(\sigma) d\sigma$$



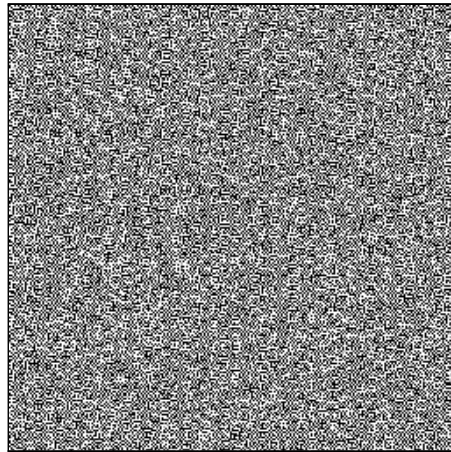
Traditional Solutions – Fourier Deconvolution

The ill-conditioning of a problem does not mean that a meaningful approximate solution cannot be computed. Rather the ill-conditioning implies that standard methods in numerical linear algebra cannot be used in a straightforward way to compute such a solution. Instead, more sophisticated methods must be applied in order to ensure the computation of a meaningful solution.

This is the essential goal of regularization methods.



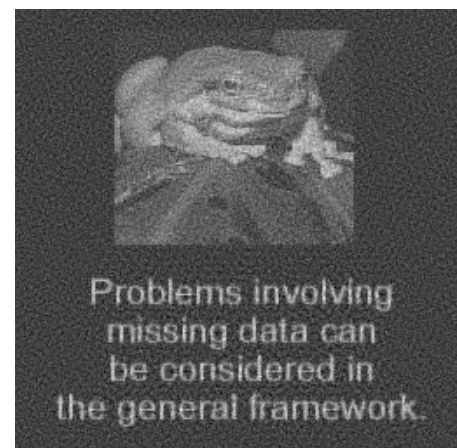
Noisey blurred image



Exact inverse

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MAP solution

Bayesian Formulation for Conductivity Imaging

	current in Ω	potential in Ω	voltage electrode	current electrode	conductivity
r.v.	R	Φ	V	J	Σ
value	ρ	ϕ	v	j	σ

Posterior

$$\begin{aligned} \Pr \{ \Sigma = \sigma, \Phi^n = \phi^n, R^n = \rho^n | \{J^n, V^n\} = \{j^n, v^n\} \} \\ = \Pr \{ \{J^n, V^n\} = \{j^n, v^n\} | \Sigma = \sigma, \Phi^n = \phi^n, R^n = \rho^n \} \\ \times \Pr \{ \Sigma = \sigma, \Phi^n = \phi^n, R^n = \rho^n \} \end{aligned}$$

$$R = -\Sigma \nabla \Phi, \phi = \Gamma_\sigma(\rho|_{\partial\Omega}) \text{ and } \rho = -\sigma \nabla \phi$$

$$\Pr \{ \Sigma = \sigma, \Phi^n = \phi^n, R^n = \rho^n \} = \Pr \{ \Sigma = \sigma, R^n = \rho^n \}$$

Stipulate $\Pr \{ \Sigma = \sigma \}$ only – usually a MRF

$$\begin{aligned} L(\sigma, \phi^n, \rho^n) \\ = \Pr \{ \{J^n, V^n\} = \{j^n, v^n\} | \Sigma = \sigma, \Phi^n = \phi^n, R^n = \rho^n \} \\ = \Pr \{ \{V^n\} = \{v^n\} | \Phi^n = \phi^n(\sigma, \rho^n) \} \\ \times \Pr \{ \{J^n\} = \{j^n\} | R^n = \rho^n \} \end{aligned}$$

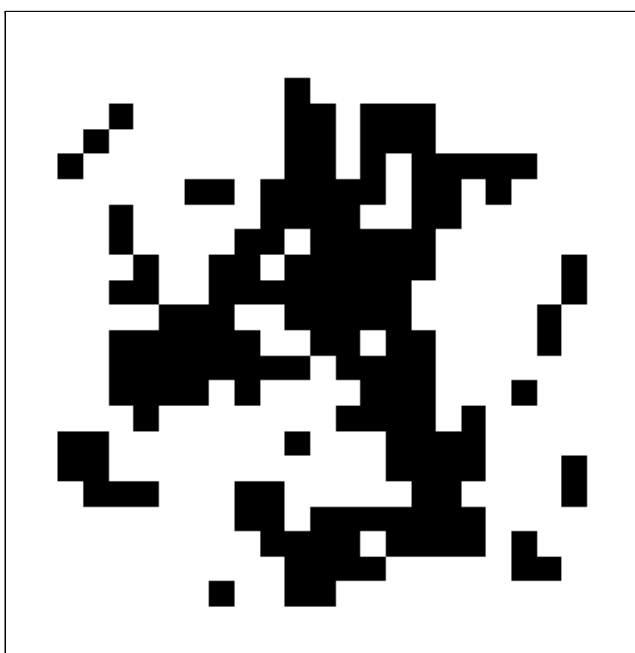
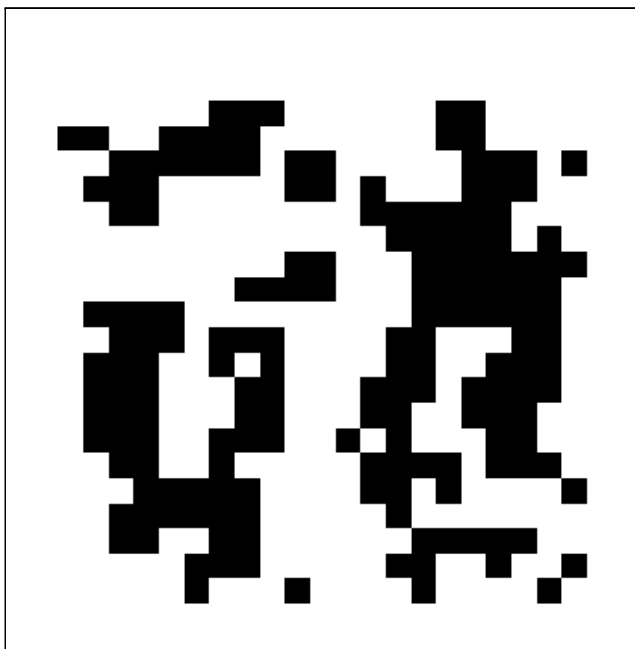
Errors i.i.d.

$$\begin{aligned} L(\sigma, \phi^n, \rho^n) &= \prod_{n=1}^N \Pr \{ V^n = v^n | \Phi^n = \Gamma_\sigma(\rho^n|_{\partial\Omega}) \} \\ &\times \Pr \{ J^n = j^n | R^n = \rho^n \}. \end{aligned}$$

Noise is normal (say)

$$\Pr \{ J = j | R = \rho \} \sim N \left((\rho(x_1), \rho(x_2), \dots, \rho(x_k))^T, s_\rho^2 \right)$$

Samples from the Prior



Markov chain Monte Carlo

- Monte Carlo integration

If $\{X_t, t = 1, 2, \dots, n\}$ are sampled from $\Pr(\sigma|v)$

$$\mathbb{E}[f(\sigma)] \approx \frac{1}{n} \sum_{t=1}^n f(X_t)$$

- Markov chain

Generate $\{X_t\}_{t=0}^{\infty}$ as a Markov chain of random variables $X_t \in \Sigma_{\Omega}$, with a t -step distribution $\Pr(X_t = \sigma | X_0 = \sigma^{(0)})$ that tends to $\Pr(\sigma|v)$, as $t \rightarrow \infty$.

Metropolis-Hastings algorithm

(1) given state σ_t at time t generate candidate state σ' from a proposal distribution $q(\cdot|\sigma_t)$

(2) Accept candidate with probability

$$\alpha(X|Y) = \min \left(1, \frac{\Pr(Y|v)q(X|Y)}{\Pr(X|v)q(Y|X)} \right)$$

(3) If accepted, $X_{t+1} = \sigma'$ otherwise $X_{t+1} = \sigma_t$

(4) Repeat

$q(\cdot|\sigma_t)$ can be any distribution that ensures the chain is:

- irreducible
- aperiodic

Three-Move Metropolis Hastings

Choose one of 3 moves with probability ζ_p , $p = 1, 2, 3$

Transition probabilities $\{\Pr^{(p)}(X_{t+1} = \sigma_{t+1} | X_t = \sigma_t)\}_{p=1}^3$
(reversible w.r.t. $\Pr(\sigma|v)$).

Overall transition probability is

$$\begin{aligned} \Pr(X_{t+1} = \sigma_{t+1} | X_t = \sigma_t) \\ = \sum_{p=1}^3 \zeta_p \Pr^{(p)}(X_{t+1} = \sigma_{t+1} | X_t = \sigma_t). \end{aligned}$$

If at least one of the moves is irreducible on Σ_Ω , then the equilibrium distribution is $\Pr(\sigma|v)$.

A pixel n is a *near-neighbour* of pixel m if their lattice distance is less than $\sqrt{8}$.

An *update-edge* is a pair of near-neighbouring pixels of unequal conductivity. $(\mathcal{N}^*(\sigma), \mathcal{N}_m^*(\sigma))$

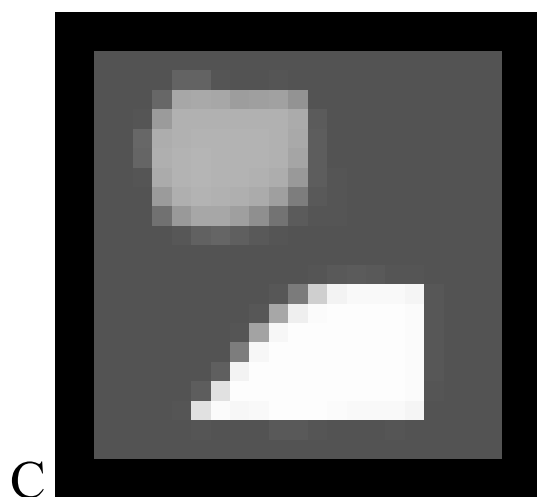
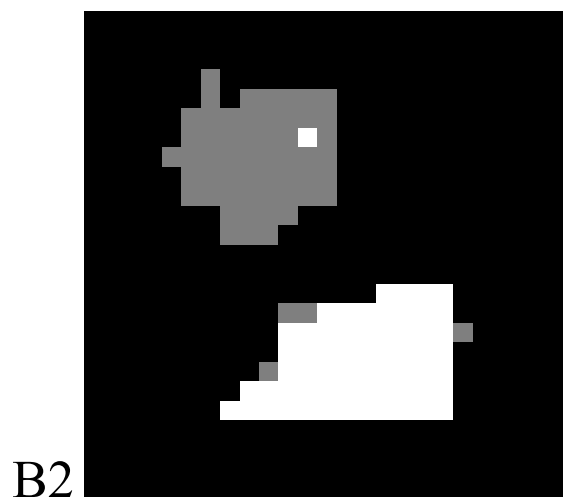
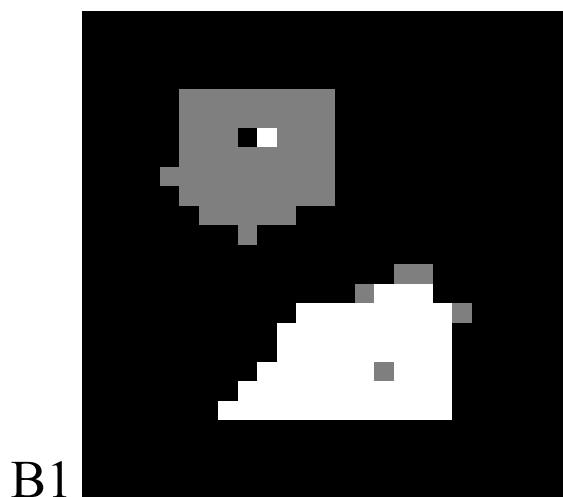
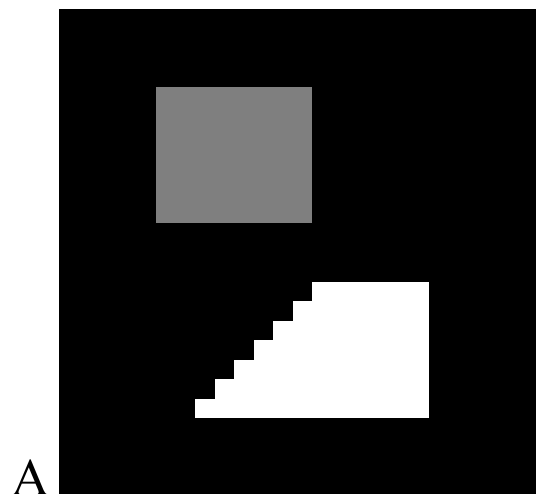
Move 1 *Flip a pixel.* Select a pixel m at random and assign σ_m a new conductivity σ'_m chosen uniformly at random from the other $\mathcal{C} - 1$ conductivity values.

Move 2 *Flip a pixel near a conductivity boundary.* Pick an update-edge at random from $\mathcal{N}^*(\sigma)$. Pick one of the two pixels in that edge at random, pixel m say. Proceed as in Move 1.

Move 3 *Swap conductivities at a pair of pixels.* Pick an update-edge at random from $\mathcal{N}^*(\sigma)$. Set $\sigma'_m = \sigma_n$ and $\sigma'_n = \sigma_m$.

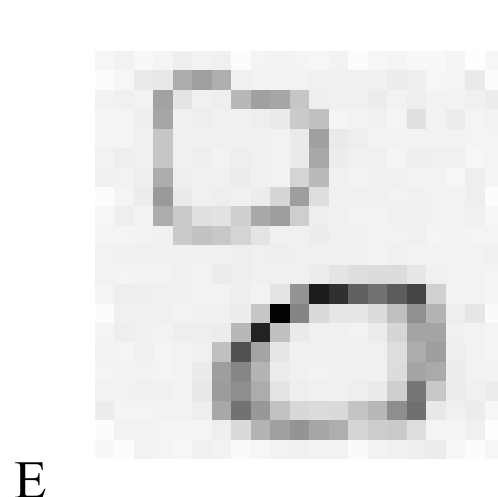
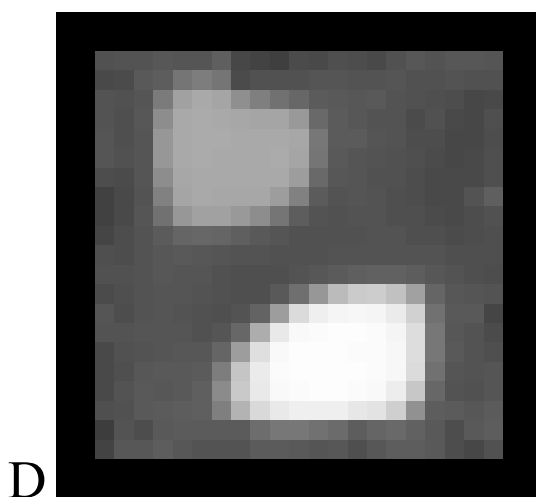
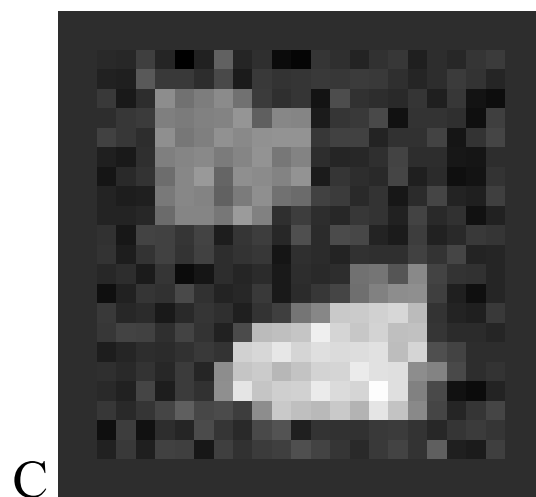
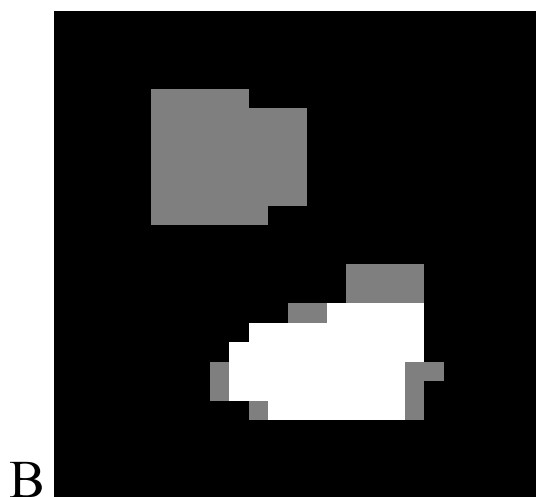
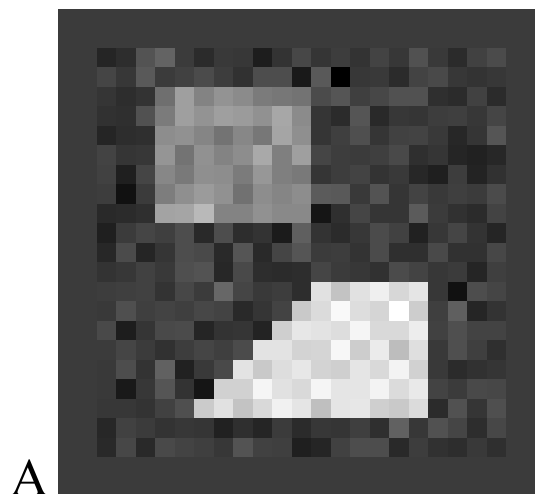
Experiment 1

(discrete variables – three conductivity levels)

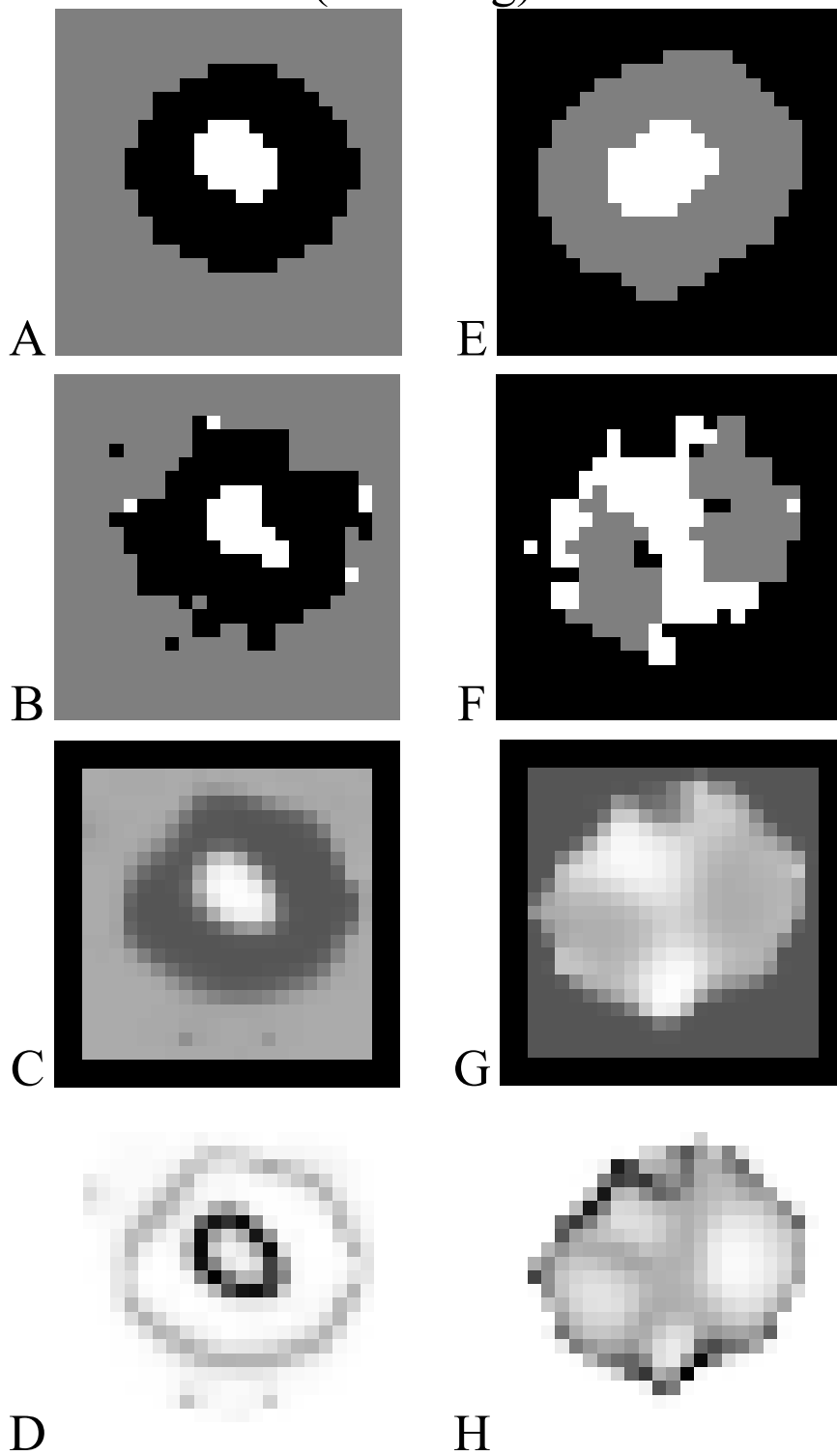


Experiment 2

(continuous variables – three conductivity types)



Experiment 3 (shielding)



Summary

- If you can simulate the forward map then you can sample and calculate expectations over the posterior, i.e., ‘solve’ the inverse problem
- Statistical inference provides a unifying framework for inverse problems